

ÉCOLE POLYTECHNIQUE FÉDÉRALE DE LAUSANNE

School of Computer and Communication Sciences

Handout 29

Information Theory and Coding

Solutions to Final Exam

Jan. 30, 2012

PROBLEM 1.

- (a) Since the events are independent $\Pr(\cap_i E_i) = \prod_i \Pr(E_i) = (1-p)^M$. From the hint $1-p \leq e^{-p}$; thus $(1-p)^M \leq \exp(-pM)$.
- (b) The decoder output is not equal to encoder input if all the events $E_i = \{\mathbf{U}(i) \neq \mathbf{u}\}$ happen. Since $\mathbf{U}(i)$ are chosen independently these events are independent and each has probability $1 - p_{U^n}(\mathbf{u})$. The claim now follows from part (a) with $p = p_{U^n}(\mathbf{u})$.
- (c) If \mathbf{u} is ϵ -typical

$$p_{U^n}(\mathbf{u}) = \prod_a p_U(a)^{\#\{i:u_i=a\}} \geq \prod_a p_U(a)^{n(1+\epsilon)p_U(a)} = 2^{-n(1+\epsilon)H(U)}.$$

The claim now follows from part (c).

- (d) Given $R > H(U)$ we can pick $\epsilon > 0$ such that $(1+\epsilon)H(U) < R$. Let T_ϵ^n denote the ϵ -typical sequences. The probability of error can be written as

$$\begin{aligned} \Pr(\text{Dec Enc}(\mathbf{U})) \neq \mathbf{U} &= \Pr(\text{Dec Enc}(\mathbf{U})) \neq \mathbf{U} | \mathbf{U} \in T_\epsilon^n) \Pr(\mathbf{U} \in T_\epsilon^n) \\ &\quad + \Pr(\text{Dec Enc}(\mathbf{U})) \neq \mathbf{U} | \mathbf{U} \notin T_\epsilon^n) \Pr(\mathbf{U} \notin T_\epsilon^n) \end{aligned}$$

The first term is less than $\Pr(\text{Dec Enc}(\mathbf{U}) \neq \mathbf{U} | \mathbf{U} \in T_\epsilon^n)$, which, in turn, is less than $\exp(-2^{n[R-(1+\epsilon)H(U)]})$ by part (c). The second term is less than $\Pr(\mathbf{U} \notin T_\epsilon^n)$. Both of these go to zero as n gets large.

PROBLEM 2.

- (a) Taking the hint, if $0 < \beta_1 < \beta_2$ we can write $\beta_1 = (1-\lambda)0 + \lambda\beta_2$ with $\lambda = \beta_1/\beta_2$. Since $\lambda \in (0, 1)$ and since f is concave

$$f(\beta_1) \geq (1-\lambda)f(0) + \lambda f(\beta_2) = \beta_1 f(\beta_2) / \beta_2$$

thus $f(\beta_1)/\beta_1 \geq f(\beta_2)/\beta_2$.

- (b) Recall that $C(\beta) = \max_{p_X: E[X] \leq \beta} I(X; Y)$. Since $E[X] = \Pr(X = 1)$, we see that $C(\beta) = \max_{\alpha \leq \beta} I(\alpha)$.

Furthermore, since $I(\cdot)$ is concave, by denoting by β_0 the value of β that maximizes $I(\beta)$, we see that $I(\cdot)$ is increasing on $[0, \beta_0]$ and decreasing on $[\beta_0, 1]$. Consequently, we can write

$$C(\beta) = \begin{cases} I(\beta) & \beta \leq \beta_0 \\ I(\beta_0) & \beta \geq \beta_0. \end{cases}$$

(c) From part (b)

$$C(\beta)/\beta = \begin{cases} I(\beta)/\beta & \beta \leq \beta_0 \\ I(\beta_0)/\beta & \beta \geq \beta_0. \end{cases}$$

By part (a) we see that $C(\beta)/\beta$ is decreasing on $[0, \beta_0]$ and continues to decrease for $\beta > \beta_0$. Consequently its supremum is achieved as β approaches 0 from above,

$$\sup_{\beta > 0} \frac{C(\beta)}{\beta} = \lim_{\beta \searrow 0} \frac{I(\beta)}{\beta}.$$

(d) Note that

$$\begin{aligned} I(\beta) &= \beta \sum_y p(y|1) \log \frac{p(y|1)}{\beta p(y|1) + (1-\beta)p(y|0)} \\ &\quad + (1-\beta) \sum_y p(y|0) \log \frac{p(y|0)}{\beta p(y|1) + (1-\beta)p(y|0)} \end{aligned}$$

and thus

$$\begin{aligned} I(\beta)/\beta &= \sum_y p(y|1) \log \frac{p(y|1)}{\beta p(y|1) + (1-\beta)p(y|0)} \\ &\quad + (1-\beta) \frac{1}{\beta} \sum_y p(y|0) \log \frac{p(y|0)}{\beta p(y|1) + (1-\beta)p(y|0)} \end{aligned}$$

As $\beta \rightarrow 0$, the first term approaches $\sum_y p(y|1) \log \frac{p(y|1)}{p(y|0)}$. So we only need to show that the second term approaches zero to reach the desired conclusion. To that end, observe that the second term, being in the form of a divergence, is non-negative. Furthermore, using $\ln z \leq z - 1$,

$$\begin{aligned} 0 &\leq \frac{1-\beta}{\beta} \sum_y p(y|0) \ln \frac{p(y|0)}{\beta p(y|1) + (1-\beta)p(y|0)} \\ &\leq \frac{1-\beta}{\beta} \sum_y p(y|0) \left[\frac{p(y|0)}{\beta p(y|1) + (1-\beta)p(y|0)} - 1 \right] \\ &= (1-\beta) \sum_y p(y|0) \frac{p(y|0) - p(y|1)}{\beta p(y|1) + (1-\beta)p(y|0)} \\ &\rightarrow \sum_y [p(y|0) - p(y|1)] = 0 \quad \text{as } \beta \rightarrow 0. \end{aligned}$$

(e) Using $[p(0|0), p(1|0)] = [1-p, p]$ and $[p(0|1), p(1|1)] = [p, 1-p]$ we see that the formula of (d) gives the capacity per cost as $(1-2p) \log[(1-p)/p]$.

PROBLEM 3.

(a) The blocklength is $n = 4$. The number of codewords is 9: we can choose any $x_3 \in \mathbb{F}_3$ and $x_4 \in \mathbb{F}_3$ and the parity check equations determine (x_1, x_2) . The rate is $R = (1/4) \log 9 = (1/2) \log 3$.

- (b) The received word \mathbf{y} is of the form $\mathbf{x} + \mathbf{z}$ where \mathbf{z} is either the zero vector or it contains a single non-zero entry. Thus $H\mathbf{y}$ is either zero or is a multiple of a column of H . Since all these cases are distinct the decoder can discover if \mathbf{x} was changed during transmission and how.
- (c) Augmenting the matrix will increase the blocklength to $n = 5$ and increase the number of codewords to 27: now (x_3, x_4, x_5) can be freely chosen. The rate thus becomes $(3/5) \log 3$ and is increased from $(1/2) \log 3$.
- (d) None of them, since each is a multiple of an existing column on H .
- (e) In this case the received word is of the form $\mathbf{y} = \mathbf{x} + \mathbf{z}$ where \mathbf{z} is either 0 or contains a single 1. To ensure $H\mathbf{y}$ is distinct in all the cases all we need to ensure is that the columns of H are non-zero and distinct (but not necessarily their multiples). Thus, $\begin{bmatrix} 2 \\ 2 \end{bmatrix}$ and $\begin{bmatrix} 2 \\ 1 \end{bmatrix}$ are valid augmentations.