

Solutions 6

1. a) Using Cauchy-Schwartz inequality, we obtain

$$|m_{2k+1}|^2 = \left| \int_{\mathbb{R}} d\mu(x) x^k x^{k+1} \right|^2 \leq \left(\int_{\mathbb{R}} d\mu(x) x^{2k} \right) \left(\int_{\mathbb{R}} d\mu(x) x^{2k+2} \right) = m_{2k} m_{2k+2},$$

so a condition on the growth of the even moments of μ ensures the same growth for the odd moments.

b) If $|m_k| \leq C^k$, then $m_{2k}^{1/2k} \leq C$, so

$$\limsup_{k \rightarrow \infty} \frac{1}{2k} (m_{2k})^{\frac{1}{2k}} = 0 < \infty.$$

If in turn the above lim sup is finite, then this means that there exists $C > 0$ such that

$$(m_{2k})^{\frac{1}{2k}} \leq C 2k, \quad \forall k \geq k_0 \text{ sufficiently large,}$$

so

$$\sum_{k=0}^{\infty} m_{2k}^{-\frac{1}{2k}} \geq \sum_{k=k_0}^{\infty} \frac{1}{C 2k} = \infty.$$

2. a) First observe that the odd moments of μ vanish, since the distribution is symmetric (i.e. $p_{\mu}(-x) = p_{\mu}(x)$ for all $x \in \mathbb{R}$). By the indicated change of variable, we have for m_{2k} :

$$\begin{aligned} m_{2k} &= \int_{-2}^2 x^{2k} \frac{1}{2\pi} \sqrt{4-x^2} dx = \frac{1}{\pi} \int_0^{\pi/2} (2 \sin(t))^{2k} \sqrt{4-4 \sin^2(t)} 2 \cos(t) dt \\ &= \frac{2^{2k+2}}{\pi} \int_0^{\pi/2} \sin(t)^{2k} \cos(t)^2 dt = \frac{2^{2k+2}}{\pi} \left(\int_0^{\pi/2} \sin(t)^{2k} dt - \int_0^{\pi/2} \sin(t)^{2(k+1)} dt \right). \end{aligned}$$

By integration by parts (with $u(t) = \sin^{2k-1}(t)$ and $v(t) = \sin(t)$), we have

$$a_{k+1} := \int_0^{\pi/2} \sin(t)^{2(k+1)} dt = (2k+1) \int_0^{\pi/2} \sin(t)^{2k} \cos^2(t) dt = (2k+1) (a_k - a_{k+1}),$$

so

$$a_{k+1} = \frac{2k+1}{2k+2} a_k = \dots = \frac{(2k+1) \cdots 3 \cdot 1}{(2k+2) \cdots 4 \cdot 2} a_0 = \frac{(2k+1)! / (2^k k!)}{2^{k+1} (k+1)!} a_0 = \frac{(2k+1)!}{2^{2k+1} k! (k+1)!} a_0.$$

Since $a_0 = \frac{\pi}{2}$, we finally obtain

$$m_{2k} = \frac{2^{2k+2}}{\pi} \frac{a_{k+1}}{2k+1} = \frac{(2k)!}{k!(k+1)!} = \frac{1}{k+1} \binom{2k}{k},$$

Since

$$m_{2k} \leq \frac{(2k)!}{(k!)^2} \leq \frac{((2k)(2k-2) \cdots 2)^2}{(k!)^2} = \frac{(2^k k!)^2}{(k!)^2} \leq 4^k,$$

Carleman's condition is satisfied. An easier way to see this is to notice that μ has bounded support $([-2, 2])$.

b) The density of $\mathcal{N}(0, 1)$ is $p_\mu(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$. Therefore $p'_\mu(x) = -x p_\mu(x)$ and we have by integration by parts:

$$\int_{\mathbb{R}} x f(x) p_\mu(x) dx = - \int_{\mathbb{R}} f(x) p'_\mu(x) dx = \int_{\mathbb{R}} f'(x) p_\mu(x) dx,$$

where we have used the fact that the boundary term vanishes to zero as $n \rightarrow \infty$ (since $f(x)$ is by assumption growing polynomially at infinity and $p_\mu(x), p'_\mu(x) \simeq \exp(-x^2/2)$).

Notice that $m_1 = 0$ and $m_2 = 1$. We then have, by application of part the preceding formula,

$$m_{k+2} = \int_{\mathbb{R}} dx p_\mu(x) x x^{k+1} = (k+1) \int_{\mathbb{R}} dx p_\mu(x) x^{k+1} = (k+1) m_k.$$

From this, we deduce by induction that $m_{2k+1} = 0$ for all $k \geq 0$ (but this could have also been deduced from the fact the μ is symmetric) and that

$$m_{2k} = (2k-1) \cdots 3 \cdot 1 = \frac{(2k)!}{2^k (k!)}, \quad \text{for all } k \geq 0.$$

This moments satisfy Carleman's condition, since as before,

$$m_{2k} \leq \frac{((2k)(2k-2) \cdots 2)^2}{2^k k!} = 2^k k!$$

and $k! \leq k^{k+1/2}$ by Stirling's formula.

c) Using the change of variable $x = e^y$, we obtain

$$m_k = \int_0^\infty x^k p_\mu(x) dx = \int_{\mathbb{R}} e^{ky} p_\mu(e^y) e^y dy.$$

Noticing that $p_\mu(e^y) e^y = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{y^2}{2}\right)$, we further obtain

$$m_k = \frac{1}{\sqrt{2\pi}} e^{k^2/2} \int_{\mathbb{R}} e^{-(y-k)^2/2} dy = e^{k^2/2}.$$

These moments do not satisfy Carleman's condition, since

$$\sum_{k=0}^{\infty} m_{2k}^{-\frac{1}{2k}} = \sum_{k=0}^{\infty} e^{-k} < \infty.$$

d) In order to satisfy $\mu(\mathbb{R}) = 1$, we must have $C = 1/\sum_{j \in \mathbb{Z}} e^{-j^2/2}$. Let us compute the moments of μ :

$$m_k = C \sum_{j \in \mathbb{Z}} e^{jk} e^{-j^2/2} = C e^{k^2/2} \sum_{j \in \mathbb{Z}} e^{-(j-k)^2/2} = e^{k^2/2}.$$

Notice that the moments of this distribution and the preceding are the same, even though the distributions are different.

e) By the change of variable $y = x^\lambda$, we have

$$\int_{\mathbb{R}} x^k d\mu(x) = c_\lambda \int_0^\infty x^k \exp(-x^\lambda) dx = c_\lambda \lambda \int_0^\infty y^{k/\lambda} e^{-y} y^{1/\lambda-1} dy = c_\lambda \lambda \Gamma((k+1)/\lambda),$$

where Γ is the Euler Gamma function. Using the approximation $\Gamma(x+1) \sim [x]!$, we see that

$$m_k = \int_{\mathbb{R}} x^k d\mu(x) \sim \left[\frac{k}{\lambda}\right]!$$

so by Stirling's formula ($\log(k!) \sim k \log k$),

$$\limsup_{k \rightarrow \infty} \frac{1}{2k} (m_{2k})^{\frac{1}{2k}} \sim \limsup_{k \rightarrow \infty} \frac{1}{2k} e^{\frac{1}{\lambda} \log(2k/\lambda)} \sim \limsup_{k \rightarrow \infty} \frac{1}{2k} (2k/\lambda)^{\frac{1}{\lambda}} < \infty$$

if and only if $\lambda \geq 1$. We can deduce the following rule of thumb from the preceding argument: a distribution is uniquely determined by its moments as long as its tail is not heavier than the exponential e^{-x} .

3. By ex. 2.b), we only need to check that for any $k \geq 0$,

$$\int_{\mathbb{R}} x^k dF_n(x) \xrightarrow{n \rightarrow \infty} m_k,$$

where $m_{2k+1} = 0$ and $m_{2k} = \frac{(2k)!}{2^k (k!)}$. Let us therefore compute

$$\int_{\mathbb{R}} x^k dF_n(x) = \mathbb{E}(X_n^k) = \frac{1}{n^{k/2}} \mathbb{E}\left((Y_1 + \dots + Y_n)^k\right)$$

Using the multinomial expansion, we obtain

$$\mathbb{E}(X_n^k) = \frac{1}{n^{k/2}} \sum_{\substack{j_1, \dots, j_n \geq 0 \\ j_1 + \dots + j_n = k}} \binom{k}{j_1, \dots, j_n} \mathbb{E}(Y_1^{j_1} \dots Y_n^{j_n}).$$

Since the Y_j are i.i.d. and $\mathbb{E}(Y_j^{2l+1}) = 0$ for all $l \geq 0$, it is easy to see that the above sum is zero if k is odd. Let us therefore consider

$$\begin{aligned} \mathbb{E}(X_n^{2k}) &= \frac{1}{n^k} \sum_{\substack{j_1, \dots, j_n \geq 0 \\ j_1 + \dots + j_n = 2k}} \binom{2k}{j_1, \dots, j_n} \mathbb{E}(Y_1^{j_1} \dots Y_n^{j_n}) \\ &= \frac{1}{n^k} \sum_{\substack{l_1, \dots, l_n \geq 0 \\ l_1 + \dots + l_n = k}} \binom{2k}{2l_1, \dots, 2l_n} \mathbb{E}(Y_1^{2l_1}) \dots \mathbb{E}(Y_n^{2l_n}). \end{aligned}$$

Let us divide this sum in two parts:

$$\begin{aligned} \mathbb{E}(X_n^{2k}) &= \frac{1}{n^k} \sum_{\substack{l_1, \dots, l_n \in \{0,1\} \\ l_1 + \dots + l_n = k}} \binom{2k}{2l_1, \dots, 2l_n} \mathbb{E}(Y_1^{2l_1}) \dots \mathbb{E}(Y_n^{2l_n}) \\ &\quad + \frac{1}{n^k} \sum_{\substack{\exists 1 \leq i \leq n : l_i \geq 2 \\ l_1 + \dots + l_n = k}} \binom{2k}{2l_1, \dots, 2l_n} \mathbb{E}(Y_1^{2l_1}) \dots \mathbb{E}(Y_n^{2l_n}). \end{aligned}$$

We see that the first term on the right-hand side is equal to

$$\frac{1}{n^k} \binom{n}{k} \frac{(2k)!}{2^k} 1 = \frac{n(n-1) \dots (n-k+1)}{n^k} \frac{(2k)!}{2^k k!} \xrightarrow{n \rightarrow \infty} \frac{(2k)!}{2^k k!}.$$

The theorem will therefore be proved if we check that the second term on the right-hand side goes to zero as $n \rightarrow \infty$:

$$\begin{aligned}
& \frac{1}{n^k} \sum_{\substack{\exists 1 \leq i \leq n: l_i \geq 2 \\ l_1 + \dots + l_n = k}} \binom{2k}{2l_1, \dots, 2l_n} \mathbb{E}(Y_1^{2l_1}) \dots \mathbb{E}(Y_n^{2l_n}) \leq \frac{C^{2k}}{n^k} \sum_{\substack{\exists 1 \leq i \leq n: l_i \geq 2 \\ l_1 + \dots + l_n = k}} \frac{(2k)!}{(2l_1)! \dots (2l_n)!} \\
& \leq \frac{C^{2k}}{n^k} n \sum_{m=2}^k \sum_{\substack{l_1, \dots, l_{n-1} \geq 0 \\ l_1 + \dots + l_{n-1} = k-m}} \frac{(2k)!}{(2l_1)! \dots (2l_{n-1})! (2m)!} \\
& \leq \frac{C^{2k} (2k)!}{n^{k-1}} \sum_{m=2}^k \sum_{\substack{l_1, \dots, l_{n-1} \geq 0 \\ l_1 + \dots + l_{n-1} = k-m}} 1 = \frac{C^{2k} (2k)!}{n^{k-1}} \sum_{m=2}^k \binom{n+k-m-2}{k-m} \\
& \leq \frac{C^{2k} (2k)!}{n^{k-1}} \sum_{m=2}^k (n+k-m-2)^{k-m} \leq \frac{C^{2k} (2k)! k}{n^{k-1}} (n+k)^{k-2} \leq \frac{C_k}{n} \xrightarrow{n \rightarrow \infty} 0. \quad \square
\end{aligned}$$