Solutions: More Exercises

Problem 1

We know that $X_i \sim \text{Bernoulli}(\theta)$ so $\mathbb{P}[X_i = 1] = \theta$ and $\mathbb{P}[X_i = 0] = 1 - \theta$.

(a) $\theta = 1/2$ so $H(X_i) = 1$ for i = 1, ..., n. By definition $(X_1, ..., X_n) = (0, ..., 0) \in A_{\epsilon}^{(n)}$ if and only if $2^{-n[H(X)+\epsilon]} \leq \mathbb{P}[(X_1, ..., X_n) = (0, ..., 0)] \leq 2^{-n[H(X)-\epsilon]},$

or

$$2^{-n[H(X)+\epsilon]} \le \prod_{i=1}^{n} \underbrace{\mathbb{P}\left[X_{i}=0\right]}_{1} \le 2^{-n[H(X)-\epsilon]}$$

Then we can take the $\log(\cdot)$ of both side of the above inequalities. Because that the $\log(\cdot)$ function is an increasing function the order of the inequalities do not change and we have

$$\begin{split} -n\left[1+\epsilon\right] &\leq -n \leq -n\left[1-\epsilon\right],\\ 1+\epsilon &\geq 1-\epsilon. \end{split}$$

So $(0, \ldots, 0) \in A_{\epsilon}^{(n)}$ if and only if $\epsilon \ge 0$ which means that it is true for every ϵ . In fact for $\theta = 1/2$ all of the 2^n possible sequences are belong to $A_{\epsilon}^{(n)}$.

(b) Again by definition $(x_1, \ldots, x_n) \in A_{\epsilon}^{(n)}$ if and only if

$$2^{-n[H(X)+\epsilon]} \le \underbrace{\mathbb{P}[(X_1=x_1]\cdots\mathbb{P}[X_n=x_n]]}_{\mathbb{P}[(X_1,\dots,X_n)=(x_1,\dots,x_n)]} \le 2^{-n[H(X)-\epsilon]},$$

if and only if

$$-n\left[H\left(X\right)+\epsilon\right] \leq \log\prod_{i=1}^{n}\mathbb{P}\left[X_{i}=x_{i}\right] \leq -n\left[H\left(X\right)-\epsilon\right],$$

if and only if

$$-n\left[H\left(X\right)+\epsilon\right] \le \log\left(\theta^{L(x^{n})}.\left(1-\theta\right)^{n-L(x^{n})}\right) \le -n\left[H\left(X\right)-\epsilon\right],$$

which only depends on $L(x^n)$.

(c) For the probability of observing a sequence $(x_1, \ldots, x_n) \in A_{\epsilon}^{(n)}$, by definition we have the following bounds

$$2^{-n[H(X)+\epsilon]} \le \mathbb{P}\left[(X_1, \dots, X_n) = (x_1, \dots, x_n) \right] \le 2^{-n[H(X)-\epsilon]},$$

so we can write

$$\frac{\mathbb{P}\left[\text{most probable sequence}\right]}{\mathbb{P}\left[\text{least probable sequence}\right]} = \frac{2^{-n[H(X)-\epsilon]}}{2^{-n[H(X)+\epsilon]}} = 2^{2n\epsilon} \stackrel{n \to \infty}{\longrightarrow} \infty,$$

which shows that the typical sequences are not "approximately equiprobable".

(d) From (b) we can write

$$-n\left[H\left(X\right)+\epsilon\right] \le L\left(x^{n}\right)\log\left(\theta\right)+\left(n-L\left(x^{n}\right)\right)\log\left(1-\theta\right) \le -n\left[H\left(X\right)-\epsilon\right],$$
$$-n\left[H\left(X\right)+\epsilon\right] \le L\left(x^{n}\right)\log\left(\frac{\theta}{1-\theta}\right)+n\log\left(1-\theta\right) \le -n\left[H\left(X\right)-\epsilon\right],$$

and

$$-n\left[H\left(X\right) + \log\left(1-\theta\right) + \epsilon\right] \le L\left(x^{n}\right)\log\left(\frac{\theta}{1-\theta}\right) \le -n\left[H\left(X\right) + \log\left(1-\theta\right) - \epsilon\right].$$

We know that

$$\begin{cases} \log\left(\frac{\theta}{1-\theta}\right) > 0 & : \quad \theta > 1/2, \\ \log\left(\frac{\theta}{1-\theta}\right) < 0 & : \quad \theta < 1/2, \end{cases}$$

so for $\theta > 1/2$ we have

$$\frac{-n\left[H(\theta) + \log\left(1 - \theta\right) + \epsilon\right]}{\log\left(\frac{\theta}{1 - \theta}\right)} \le L\left(x^n\right) \le \frac{-n\left[H(\theta) + \log\left(1 - \theta\right) - \epsilon\right]}{\log\left(\frac{\theta}{1 - \theta}\right)}$$

where $H(\theta) = -\theta \log (\theta) - (1 - \theta) \log (1 - \theta)$. Then we have

$$n \cdot \theta - n \cdot \frac{\epsilon}{\log\left(\frac{\theta}{1-\theta}\right)} \le L\left(x^n\right) \le n \cdot \theta + n \cdot \frac{\epsilon}{\log\left(\frac{\theta}{1-\theta}\right)}$$

So by choosing $p = \theta$ and $\alpha = \frac{\epsilon}{\log\left(\frac{\theta}{1-\theta}\right)}$ we have $C^{(n)}(\alpha, p) = A_{\epsilon}^{(n)}$.

Problem 2

Every set S_i is chosen uniformly at random from all 2^m possible subsets of the set $\{1, \ldots, m\}$. Then at step *i* we ask the question: is $X \in S_i$? Based on the answers to these questions we want to find the value of X.

(a) Let us assume that X = 1. Then we can write the event that the question S_i at *i*th step has the same answer for object 1 and 2 as follows

same answer at *i*th step =
$$[1 \in S_i, 2 \in S_i] \cup [1 \notin S_i, 2 \notin S_i]$$

because $A_i \cap B_i = \emptyset$. So for the probabilities we have

 $\mathbb{P}\left[\text{same answer at } i\text{th step}\right] = \mathbb{P}\left[1 \in S_i, 2 \in S_i\right] + \mathbb{P}\left[1 \notin S_i, 2 \notin S_i\right] = \mathbb{P}\left[A_i\right] + \mathbb{P}\left[B_i\right]$

Now we can use the chain rule to write

$$\mathbb{P}\left[1 \in S_i, 2 \in S_i\right] = \mathbb{P}\left[2 \in S_i \mid 1 \in S_i\right] \mathbb{P}\left[1 \in S_i\right].$$

Then it can be easily observed that

$$\mathbb{P}\left[1 \in S_i\right] = \frac{2^{m-1}}{2^m} = 1/2,$$

and

$$\mathbb{P}\left[2 \in S_i \mid 1 \in S_i\right] = \frac{2^{m-2}}{2^{m-1}} = 1/2,$$

 \mathbf{SO}

$$\mathbb{P}[A_i] = \mathbb{P}[1 \in S_i, 2 \in S_i] = \frac{1}{4}.$$

Using a similar argument we can show that

$$\mathbb{P}[B_i] = \mathbb{P}[1 \notin S_i, 2 \notin S_i] = \frac{1}{4}$$

So we have

$$\mathbb{P}[\text{same answer with object 2 after } k \text{ step}] = \prod_{i=1}^{k} \mathbb{P}[A_i \cup B_i] = \prod_{i=1}^{k} \mathbb{P}[A_i] + \mathbb{P}[B_i] = 1/2^k.$$

(b) We can use the argument of part (a) to write

 $\mathbb{P}[\text{same answer with object } i \text{ after } k \text{ step}] = 1/2^k,$

so the expected number of objects in the set $\{2, \ldots, m\}$ that have the same answers to the k questions as does the correct object 1 is

average number of wrong objects
$$=\frac{m-1}{2^k}=\frac{2^n-1}{2^k}=2^{n-k}-2^{-k}.$$

(c) Choosing
$$k = n + \sqrt{n}$$
 we have

average number of wrong objects = $2^{n-n-\sqrt{n}} - 2^{-n-\sqrt{n}} = 2^{-\sqrt{n}} - 2^{-n-\sqrt{n}}$.

(d) Let *E* be the number of wrong objects after asking *k* questions. This is a random variable that in part (b) and (c) we calculated its expected value. For $k = n + \sqrt{n}$ we found that $\mathbb{E}[E] = 2^{-\sqrt{n}} - 2^{-n-\sqrt{n}}$. Now we have

$$\lim_{n \to \infty} \mathbb{E}[E] \to 0.$$

But we want to show that the probability of wrong answers goes to zero as $n \to \infty$. To this end, we use the Markov's inequality as following

$$\mathbb{P}\left[E \ge 1\right] = \mathbb{P}\left[E \ge \frac{1}{\underbrace{2^{-\sqrt{n}} - 2^{-n+\sqrt{n}}}_{t}} \underbrace{(2^{-\sqrt{n}} - 2^{-n-\sqrt{n}})}_{\mu}\right] \stackrel{(i)}{\le} 2^{-\sqrt{n}} - 2^{-n-\sqrt{n}},$$

where (i) follows from Markov's inequality. So for the probability of having wrong answer we have

$$\lim_{k \to \infty} \mathbb{P}\left[E \ge 1\right] \longrightarrow 0.$$

Because the random variable E takes its value from $\{0, 1, 2, ...\}$ we conclude that E = 0 with probability approaching 1 as $n \to \infty$.

Problem 3

 \mathcal{P} is a given set of participants and \mathcal{A} is a collection of subsets of \mathcal{P} .

(a) We know that $A, B \subset \mathcal{P}, B \notin \mathcal{A}$ and $A \cup B \in \mathcal{A}$. Using chain rule we can write

$$H(X_A, S \mid X_B) = H(X_A \mid X_B) + \underbrace{H(S \mid X_A, X_B)}_{0},$$

so we have $H(X_A \mid X_B) = H(X_A, S \mid X_B)$. Again applying the chain rule we obtain

$$H(X_A \mid X_B) = H(X_A, S \mid X_B) = \underbrace{H(S \mid X_B)}_{H(S)} + H(X_A \mid X_B, S),$$

and we are done.

(b) In this part we have $B \in \mathcal{A}$. Using the chain rule we can write

$$H(X_{A} | X_{B}) = H(X_{A}, S | X_{B}) - H(S | X_{A}, X_{B})$$

= $H(X_{A} | X_{B}, S) + \underbrace{H(S | X_{B})}_{0} - \underbrace{H(S | X_{A}, X_{B})}_{0}$

The last term is zero because we have $H(S \mid X_A, X_B) \leq H(S \mid X_B) = 0$, since conditioning reduces the entropy.

(c) Here we assume that $A, B, C \subset \mathcal{P}$ where $A \cup C \in \mathcal{A}$, $B \cup C \in \mathcal{A}$, and $C \notin \mathcal{A}$. Then we can write

$$I(X_A; X_B \mid X_C) = \underbrace{H\left(X_B \mid \underbrace{X_C}_{\notin A}\right)}_{H(S) + H(X_B \mid X_C, S)} - \underbrace{H\left(X_B \mid \underbrace{X_A, X_C}_{\in A}\right)}_{H(X_B \mid X_A, X_C, S)}$$
$$= H(S) + \underbrace{I(X_B; X_A \mid X_C, S)}_{>0},$$

so we have $I(X_A; X_B \mid X_C) \ge H(S)$.

Problem 4

Two random variables X and X' are i.i.d. with entropy H(X).

(a) We want to show that $\mathbb{P}[X = X'] \ge 2^{-H(X)}$. Suppose that $X \sim P(x)$ and let us write

$$2^{-H(X)} = 2^{\mathbb{E}[\log P(X)]}$$

$$\stackrel{(i)}{\leq} \mathbb{E}\left[2^{\log P(X)}\right]$$

$$= \sum_{x} P(x) 2^{\log P(x)}$$

$$= \sum_{x} P(x)^{2}$$

$$= \mathbb{P}\left[X = X'\right],$$

where (i) comes from the Jensen's inequality applying on the function $f(y) = 2^y$. Because the function f is convex we have $\mathbb{E}f(Y) \ge f(\mathbb{E}Y)$. So defining a new random variable $Y \triangleq \log P(X)$ we have

$$2^{\mathbb{E}\log P(X)} \leq \mathbb{E}^{2^{\mathbb{E}\log P(X)}},$$

which results in (i).

(b) This part is also very similar to the previous part. Let us write

$$2^{-H(P)-D(P||Q)} = 2^{\sum P(x)\log P(x) + \sum P(x)\log \frac{Q(x)}{P(x)}} = 2^{\sum P(x)\log Q(x)}$$
$$= 2^{\mathbb{E}_p \log Q(x)} \leq \mathbb{E}_p 2^{\log Q(x)} = \sum P(x) 2^{\log Q(x)}$$
$$= \sum P(x) Q(x)$$
$$= \mathbb{P} [X = X'].$$

The same method applies for the other one.

Problem 5

(a) Suppose that we have two stochastic processes X_1, \ldots, X_n and Y_1, \ldots, Y_n such that $Y_i = \Phi(X_i)$ for $i = 1, 2, \ldots$, where $\Phi(\cdot)$ is some deterministic function. Then we can write

$$(Y_1,\ldots,Y_n)=(\Phi(X_1),\ldots,\Phi(X_n))=F(X_1,\ldots,X_n),$$

or

$$\left(Y_{1}^{n}\right)=F\left(X_{1}^{n}\right),$$

for some deterministic multivariate function $F(\cdot)$. Now using the chain rule we have

$$H(X_{1}^{n}, Y_{1}^{n}) = \begin{cases} H(X_{1}^{n}) + \overbrace{H(Y_{1}^{n} \mid X_{1}^{n})}^{=0}, \\ \sum_{i=0}^{20} \\ H(Y_{1}^{n}) + \overbrace{H(X_{1}^{n} \mid Y_{1}^{n})}^{=0}, \end{cases}$$

so we can conclude that $H(X_1^n) \ge H(Y_1^n)$. Then we take the limit as follows

$$\lim_{n \to \infty} \frac{1}{n} H\left(X_1^n\right) \ge \lim_{n \to \infty} \frac{1}{n} H\left(Y_1^n\right),$$

which results in

$$H\left(\mathcal{Y}\right) \leq H\left(\mathcal{X}\right).$$

- (b) Refer to the proof of part (c).
- (c) $Z_i = \Psi(X_i, \ldots, X_{i+l})$ for $i = 1, 2, \ldots$, and $i \le l \le n$ where l is a fixed number. Then we can write

$$(Z_1, \ldots, Z_{n-l}) = [\Psi(X_1, \ldots, X_{1+l}), \ldots, \Psi(X_{n-l}, \ldots, X_n)] = F(X_1, \ldots, X_n),$$

where $F(\cdot)$ is a deterministic multivariate function. Using the same argument given in part (a) we have

$$H\left(Z_1^{n-l}\right) \le H\left(X_1^n\right).$$

Multiplying both side by 1/n and taking the limit we obtain

$$\lim_{n \to \infty} \frac{n-l}{n} \frac{1}{n-l} H\left(Z_1^{n-l}\right) \le \lim_{n \to \infty} \frac{1}{n} H\left(X_1^n\right),$$

or we can write

$$\lim_{n \to \infty} \frac{1}{n-l} H\left(Z_1^{n-l}\right) \le \lim_{n \to \infty} \frac{1}{n} H\left(X_1^n\right),$$

because for fixed values of l we have $\lim_{n\to\infty} \frac{n-l}{n} = 1$. Then we can conclude that

$$H\left(\mathcal{Z}\right) \leq H\left(\mathcal{Y}\right).$$

(d) We have a second order Markov process over alphabet $\{0.1\}$ which realized as follows

$$X_n = \begin{cases} \text{Bernoulli} (0.5) & \text{if } X_{n-1} = X_{n-2}, \\ \\ \text{Bernoulli} (0.9) & \text{if } X_{n-1} \neq X_{n-2}. \end{cases}$$

Let us define the Markov process $Z_n = (X_n, X_{n+1})$. Obviously Z_n is a first order Markov process or a Markov chain (it is a process that the current state only depends on the last state). From the definition of process Z_n we can write

$$H(Z_1,\ldots,Z_n)=H(X_1,\ldots,X_{n+1}).$$

Taking the limit we can write

$$\lim_{n \to \infty} \frac{1}{n} H\left(Z_1, \dots, Z_n\right) = \lim_{n \to \infty} \frac{n+1}{n} \frac{1}{n+1} H\left(X_1, \dots, X_{n+1}\right)$$

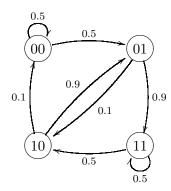
so we have

$$H\left(\mathcal{Z}\right) = H\left(\mathcal{X}\right).$$

To find the entropy rate of process X_n we can find the entropy rate of process Z_n that is a first order Markov chain with the states: 00,01,10,11. From the definition of the process X_n we can derive the transition matrix of the Markov chain Z_n which is as follows

$$P = \begin{array}{cccc} 00 & 01 & 10 & 11 \\ 00 & 0.5 & 0.5 & 0 & 0 \\ 01 & 0.5 & 0.5 & 0 & 0 \\ 0 & 0 & 0.1 & 0.9 \\ 0.1 & 0.9 & 0 & 0 \\ 0 & 0 & 0.5 & 0.5 \end{array} \right],$$

which is equivalent to the following transition graph



To find the entropy rate of the Markov chain Z_n firstly we have to find it stationary distribution μ which is the solution of the following system of equations

$$\mu = \mu P.$$

Solving the above system of equations we obtain

Then for the entropy rates we can write

$$H(\mathcal{X}) = H(\mathcal{Z}) = H(Z_2 \mid Z_1) = \sum_{i \in \{00,01,10,11\}} \mu_i \sum_{j \in \{00,01,10,11\}} P_{ij} \log P_{ij}.$$