

Solutions: Homework Set # 10

Problem 1 (A MUTUAL INFORMATION GAME)

- (a) Let $a_b = \arg \min_a f(a, b)$, *i.e.*, $f(a, b) \geq f(a_b, b), \forall a, b$. Then by taking the maximum of both sides,

$$\max_b f(a, b) \geq \max_b f(a_b, b) = \max_b \min_a f(a, b).$$

Note that the RHS does not depend on a anymore, while the LHS still depends on a . However, since the inequality holds for all a and b , it also holds for the minimizing a , *i.e.*,

$$\min_a \max_b f(a, b) \geq \max_b \min_a f(a, b).$$

- (b)

$$\begin{aligned} I(X; X + Z^*) &= h(X + Z^*) - h(X + Z^*|X) \\ &= h(X + Z^*) - h(Z^*) \\ &\leq h(X^* + Z^*) - h(Z^*) \\ &= I(X^*; X^* + Z^*) \end{aligned}$$

where the inequality follows from the fact that given the variance, the entropy is maximized by the Gaussian distribution.

- (c) 1. This is just expansion of mutual information as $I(X; X + Z) = h(X + Z) - h(X + Z|Z) = h(Y) - h(Z)$ since X and Z are independent.
 2. Each entropy expression is replaced by its definition.
 3. Note that $f_{Y^*}(y) = \frac{1}{\sqrt{2\pi(P+N)}} \exp\left(-\frac{y^2}{2(P+N)}\right)$. Therefore, $\log f_{Y^*}(y) = -\frac{y^2}{2(P+N)} - \frac{1}{2} \log 2\pi(P+N)$.

$$\begin{aligned} \int_y f_{Y^*}(y) \log f_{Y^*}(y) dy &= \int_y f_{Y^*}(y) \left[-\frac{y^2}{2(P+N)} - \frac{1}{2} \log 2\pi(P+N) \right] dy \\ &= -\frac{1}{2(P+N)} \int_y y^2 f_{Y^*}(y) dy - \frac{1}{2} \log 2\pi(P+N) \int_y f_{Y^*}(y) dy \\ &= -\frac{1}{2(P+N)} \mathbb{E}_{Y^*}[y^2] - \frac{1}{2} \log 2\pi(P+N) \\ &\stackrel{(a)}{=} -\frac{1}{2(P+N)} \mathbb{E}_Y[y^2] - \frac{1}{2} \log 2\pi(P+N) \\ &= -\frac{1}{2(P+N)} \int_y y^2 f_Y(y) dy - \frac{1}{2} \log 2\pi(P+N) \int_y f_Y(y) dy \\ &= \int_y f_Y(y) \left[-\frac{y^2}{2(P+N)} - \frac{1}{2} \log 2\pi(P+N) \right] dy \\ &= \int_y f_Y(y) \log f_{Y^*}(y) dy \end{aligned}$$

where (a) follows from the fact that $\mathbb{E}_{Y^*}[y^2] = \mathbb{E}_Y[y^2]$. The same proof holds for Z and Z^* .

4. Integration is a linear operation, and

$$\begin{aligned} - \int_y f_{Y^*}(y) \log f_{Y^*}(y) dy + \int_y f_Y(y) \log f_Y(y) dy &= \int_y f_Y(y) [\log f_Y(y) - \log f_{Y^*}(y)] dy \\ &= \int_y f_Y(y) \log \frac{f_Y(y)}{f_{Y^*}(y)} dy. \end{aligned}$$

Similarly, we can rewrite the two integrals on Z .

5.

$$\begin{aligned} \int_y f_Y(y) \log \frac{f_Y(y)}{f_{Y^*}(y)} dy + \int_z f_Z(z) \log \frac{f_{Z^*}(z)}{f_Z(z)} dz \\ &= \int_y \left(\int_z f_{Y,Z}(y, z) dz \right) \log \frac{f_Y(y)}{f_{Y^*}(y)} dy + \int_z \left(\int_y f_{Y,Z}(y, z) dy \right) \log \frac{f_{Z^*}(z)}{f_Z(z)} dz dy \\ &= \int_y \int_z f_{Y,Z}(y, z) \log \frac{f_Y(y)}{f_{Y^*}(y)} dz dy + \int_y \int_z f_{Y,Z}(y, z) \log \frac{f_{Z^*}(z)}{f_Z(z)} dz dy \\ &= \int_y \int_z f_{Y,Z}(y, z) \log \frac{f_Y(y) f_{Z^*}(z)}{f_{Y^*}(y) f_Z(z)} dz dy \end{aligned}$$

6. By concavity of the function $\log(\cdot)$.

7. Note that $Y = X^* + Z$. Therefore

$$f_{Y,Z}(y, z) = f_{X^*,Z}(y - z, z) = f_{X^*}(y - z) f_Z(z),$$

where the last equality follows from the fact that X^* and Z are independent.

8. This should be in fact equality. The reason is that $f_Z(z)$ can be cancelled from the nominator and the denominator, and then we take out every term does not depend on z from the inner integral.

9. Again since $Y^* = X^* + Z^*$, and X^* and Z^* are independent, we have $f_{Y^*}(y) = \int_z f_{X^*}(y - z) f_{Z^*}(z) dz$.

10. By cancelling $f^{Y^*}(y)$, the remaining would be $\int_y f_Y(y) dy$ which equals 1 since $f_Y(y)$ is a probability distribution.

(d) Using parts (b) and (c) we have,

$$\begin{aligned} \min_{p(z)} \max_{p(x)} I(X; X + Z) &\leq \max_{p(x)} \leq \max_{p(x)} I(X; X + Z^*) \\ &= I(X^*; X^* + Z^*) \\ &= \min_{p(z)} I(X^*; X^* + Z) \\ &\leq \max_{p(x)} \min_{p(z)} I(X; X + Z).. \end{aligned} \tag{1}$$

On the other hand, the result of part (a) for $f(p(z), p(x)) = I(X; X + Z)$ gives us

$$\min_{p(z)} \max_{p(x)} I(X; X + Z) \geq \max_{p(x)} \min_{p(z)} I(X; X + Z). \tag{2}$$

Combining (1) and (2), we have

$$\begin{aligned} \min_{p(z) \in \mathbb{F}_N} \max_{p(x) \in \mathbb{F}_P} I(X; X + Z) &= \max_{p(x) \in \mathbb{F}_P} \min_{p(z) \in \mathbb{F}_P} I(X; X + Z) \\ &= I(X^*; X^* + Z^*) \\ &= \frac{1}{2} \log \left(1 + \frac{P}{N} \right). \end{aligned}$$

Problem 2 (ERASURE DISTORTION)

The rate distortion function is given by

$$R(D) = \min_{p(\hat{x}|x): \sum p(x, \hat{x}) d(x, \hat{x}) \leq D} I(X; \hat{X}),$$

we proceed by finding the minimizing $p(\hat{x}|x)$. The infinite distortion constrains $p(0|1) = p(1|0) = 0$. By symmetry, $p(E|0) = p(E|1) = \alpha$ and $p(0|0) = p(1|1) = 1 - \alpha$.

For this distribution the distortion is $\sum p(x, \hat{x}) d(x, \hat{x}) = \alpha \leq D$ and $I(X; \hat{X}) = 1 - \alpha$ which is minimized for $D = \alpha$. So the rate distortion function is $R(D) = 1 - D$ for $0 \leq D \leq 1$, and $R(D) = 0$ for $D > 1$.

To achieve this rate distortion function, we can proceed as follows: if D is rational (e.g. $D = \frac{k}{n}$) then we send only $n - k$ of any block of n bits. We reproduce these bits exactly and reproduce the remaining bits as erasures. Hence we can send information at rate $1 - D$ and achieve a distortion D . If D is irrational, we can get arbitrarily close to D by using longer and longer block lengths.

Problem 3 (CONVEXITY OF MUTUAL INFORMATION AS A FUNCTION OF $w(y|x)$)

Consider the following chain of inequalities and equalities:

$$\begin{aligned} I(X; Y_\lambda | Z) &= h(X|Z) - h(X|Y_\lambda, Z) \\ &\stackrel{(a)}{=} h(X) - h(X|Y_\lambda, Z) \\ &\stackrel{(b)}{\geq} h(X) - h(X|Y_\lambda) \\ &= I(X; Y_\lambda) \end{aligned}$$

where (a) follows from independence of X and Z , and (b) follows since conditioning cannot increase entropy ($-h(X|Y, Z) \geq -h(X|Y)$). Also, notice that

$$\begin{aligned} I(X; Y_\lambda | Z) &= I(X; Y_\lambda | Z = 1) Pr(Z = 1) + I(X; Y_\lambda | Z = 2) Pr(Z = 2) \\ &= I(X; Y_1) \lambda + I(X; Y_2) (1 - \lambda). \end{aligned}$$

Hence, $I(X; Y_1) \lambda + I(X; Y_2) (1 - \lambda) \geq I(X; Y_\lambda)$, i.e. mutual information is convex in $w(y|x)$.

Problem 4

(a)

$$I(X^m; \hat{Y}^m) = h(X^m) - h(X^m | \hat{X}^m) \quad (3)$$

$$= \sum_{i=1}^m h(X_i) - \sum_{i=1}^m h(X_i | X^{i-1}, \hat{X}^m) \quad (4)$$

$$\geq \sum_i h(X_i) - \sum_i h(X_i | \hat{X}_i) \quad \text{Follows by the fact that conditioning reduces entropy} \quad (5)$$

$$= \sum_i I(X_i; \hat{X}_i) \quad (6)$$

$$\geq \sum_i \min_{f(\hat{x}_i|x_i): \mathbb{E}(\hat{X}_i - X_i)^2 \leq D_i} I(X_i; \hat{X}_i) \quad (7)$$

$$= R(D_i) \quad \text{by definition} \quad (8)$$

$$= \sum_{i=1}^m \left(\frac{1}{2} \log \frac{\sigma_i^2}{D_i} \right)^+ . \quad \text{As you have derived in class, where } D_i = \mathbb{E}(\hat{X}_i - X_i)^2. \quad (9)$$

(b) (5) is tight when $f(x^m | \hat{x}^m) = \prod_{i=1}^m f(x_i | \hat{x}_i)$.

(7) is tight for $f(x_i | \hat{x}_i)$ that achieves $I(X_i; \hat{X}_i) = \left(\frac{1}{2} \log \frac{\sigma_i^2}{D_i} \right)^+$. For $D_i \leq \sigma_i$, choose $f(\hat{x}_i, x_i)$ such that $X_i = \hat{X}_i + Z_i$, \hat{X}_i, Z_i being independent and each of $\hat{X}_i \sim \mathcal{N}(0, \sigma_i^2 - D_i)$ and $Z_i \sim \mathcal{N}(0, D_i)$. Thus $I(X_i; \hat{X}_i) = \frac{1}{2} \log \frac{\sigma_i^2}{D_i}$. for $D_i \geq \sigma_i$, choose $\hat{X}_i = 0$ with probability 1.

(c) So finally,

$$R(D) = \min_{f(\hat{x}^m|x^m): \mathbb{E}d(X^m, \hat{X}^m) \leq D} I(X^m; \hat{X}^m),$$

where $\mathbb{E}d(X^m, \hat{X}^m) = \mathbb{E} \sum_i (\hat{X}_i - X_i)^2 = \sum_i \mathbb{E}(\hat{X}_i - X_i)^2$.

For $f(x^m | \hat{x}^m)$ that you found in part(b),

$$I(X^m; \hat{Y}^m) = \sum_{i=1}^m \left(\frac{1}{2} \log \frac{\sigma_i^2}{D_i} \right)^+,$$

with $D_i = \mathbb{E}(\hat{X}_i - X_i)^2$. Thus the rate distortion function can be reduced to the following optimization problem:

$$R(D) = \min_{\sum_i D_i \leq D} \sum_{i=1}^m \left(\frac{1}{2} \log \frac{\sigma_i^2}{D_i} \right)^+ .$$

(d) Let us work with \ln rather than \log in this part. Using Lagrange multipliers, we construct the functional

$$J(D) = \sum_{i=1}^m \frac{1}{2} \ln \frac{\sigma_i^2}{D_i} + \lambda \sum_{i=1}^m D_i.$$

Differentiating with respect to D_i and setting equal to 0,

$$\frac{\partial J}{\partial D_i} = -\frac{1}{2D_i} + \lambda = 0 \implies D_i = \lambda'.$$

Thus the optimum allotment of bits to the various descriptions results in an equal distortion for each random variable as long as $\lambda' \leq \sigma_i^2$ for all i . As the total allowable

distortion D increases, the constant λ' increases until it exceeds σ_i^2 for some i . At this point, the optimum solution is on the boundary of the allowable region of distortion and Kuhn-Tucker conditions should give the answer:

λ is chosen such that

$$\frac{\partial J}{\partial D_i} = -\frac{1}{2D_i} + \lambda \begin{cases} = 0 & \text{if } D_i < \sigma_i^2 \\ \leq 0 & \text{if } D_i \geq \sigma_i^2 \end{cases}$$

Now we check that

$$D_i = \begin{cases} \lambda' & \text{if } \lambda' < \sigma_i^2 \\ \sigma_i^2 & \text{if } \lambda' \geq \sigma_i^2 \end{cases}$$

in fact satisfy K-T conditions: For $D_i = \lambda' \leq \sigma_i^2$, $-\frac{1}{2D_i} + \lambda = 0$ as required. For $D_i = \sigma_i^2 \leq \lambda'$, $-\frac{1}{2D_i} + \lambda \leq -\frac{1}{2\lambda'} + \lambda \leq 0$ as required again. λ' is chosen so that $\sum_i D_i = D$. By abuse of notation, λ' is λ of part (d) of the homework sheet.